

Using the package `glmBfp`: a binary regression example.

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This short vignette shall introduce into the usage of the package `glmBfp`. For more information on the methodology, see Sabanés Bové and Held (2011). There you can also find the references for the other tools mentioned here.

If you have any questions or critique concerning the package, write an email to me: `daniel.sabanesbove@ifspm.uzh.ch`. Many thanks in advance!

0.1 Pima Indians diabetes data

We will have a look at the Pima Indians diabetes data, which is available in the package `MASS`:

```
> library(MASS)
> pima <- rbind(Pima.tr, Pima.te)
> pima$hasDiabetes <- as.numeric(pima$type == "Yes")
> pima.nObs <- nrow(pima)
```

Setup For $n = 532$ women of Pima Indian heritage, seven possible predictors for the presence of diabetes are recorded. We would like to investigate with a binary regression, which of them are relevant, and what form the statistical association has – is it a linear effect, or rather a nonlinear effect? Here we will model possible nonlinear effects with the fractional polynomials.

First, we need to decide on the prior distributions to use. We are going to use the generalised hyper- g priors for GLMs (Sabanés Bové and Held, 2011). They are automatic and are supposed to yield reasonable results. We only need to specify which hyper-prior to put on the factor g . One possible choice is the Zellner-Siow hyper-prior which says $g \sim \text{IG}(1/2, n/2)$:

```
> library(glmBfp)
> ## define the prior distributions for g which we are going to use:
> prior <- InvGammaGPrior(a=1/2,
+                          b=pima.nObs/2)
> ## (the warning can be ignored)
```

This corresponds to the F1 prior in Sabanés Bové and Held (2011). Another possible choice is the hyper- g/n prior. For this there is no special constructor function, instead you can directly specify the log prior density, as follows:

```
> ## You may also use the hyper-g/n prior:
> prior.f2 <- CustomGPrior(logDens=function(g)
+                          - log(pima.nObs) - 2 * log(1 + g / pima.nObs))
```

Stochastic model search Next, we will do a stochastic search on the (very large) model space to find “good” models. Here we have to decide on the model prior, and in this example we use the `sparse` type which was also used in the paper. We use a `chainlength`

of 100, which is very small but enough for illustration purposes (usually one should use at least 10 000 as a rule of thumb), and save all models (in general `nModels` is the number of models which are saved from all visited models). Finally, we decide that we do not want to use OpenMP acceleration (this would parallelise loops over all observations on all cores of your processor) and that we want to do higher order correction for the Laplace approximations. In order to be able to reproduce the analysis, it is advisable to set a seed for the random number generator before starting the stochastic search.

```
> set.seed(102)
> time.pima <-
+   system.time(models.pima <-
+     glmBayesMfp(type ~
+       bfp(npreg) +
+       bfp(glu) +
+       bfp(bp) +
+       bfp(skin) +
+       bfp(bmi) +
+       bfp(ped) +
+       bfp(age),
+       data=pima,
+       family=binomial("logit"),
+       priorSpecs=
+       list(gPrior=prior,
+           modelPrior="sparse"),
+       nModels=1e3L,
+       chainlength=1e1L,
+       method="sampling",
+       useOpenMP=FALSE,
+       higherOrderCorrection=TRUE))
```

Starting sampler...

```
0%-----
-----
Number of non-identifiable model proposals:    0
Number of total cached models:                 9
Number of returned models:                     9
```

```
> time.pima
```

```
   user  system elapsed
0.376   0.000   0.376
```

```
> attr(models.pima, "numVisited")
```

```
[1] 9
```

We see that the search took 0 seconds, and 9 models were found. Now, if we want to have a table of the found models, with their posterior probability, the log marginal likelihood, the log prior probability, and the powers for every covariate and the number of times that the sampler encountered that model:

```
> table.pima <- as.data.frame(models.pima)
> table.pima
```

	posterior	logMargLik	logPrior	age	bmi	bp	glu	npreg	ped	skin
1	5.752526e-01	-265.8779	-11.849169					1		-0.5
2	4.036782e-01	-266.2321	-11.849169					2		-0.5
3	7.971175e-03	-270.1569	-11.849169					2		3
4	6.638447e-03	-268.2604	-13.928611			0	2			-0.5
5	6.134410e-03	-272.4983	-9.769728					2		
6	3.250886e-04	-273.3564	-11.849169			3	2			
7	2.089908e-20	-310.6396	-11.849169	2						-0.5
8	2.131992e-23	-317.5274	-11.849169						2	3
9	2.726801e-31	-339.8609	-7.690286							

Note that while `frequency` refers to the frequency of the models in the sampling chain, thus providing a Monte Carlo estimate of the posterior model probabilities, `posterior` refers to the renormalised posterior model probabilities. The latter has the advantage that ratios of posterior probabilities between any two models are exact, while the former is unbiased (but obviously has larger variance).

Inclusion probabilities The estimated marginal inclusion probabilities for all covariates are also saved:

```
> round(attr(models.pima, "inclusionProbs"),2)
```

age	bmi	bp	glu	npreg	ped	skin
0.00	0.00	0.01	1.00	0.00	0.00	0.99

Sampling model parameters If we now want to look at the estimated covariate effects in the estimated MAP model which has the configuration given in the last seven columns of `table.pima`, then we first need to generate parameter samples from that model:

```
> ## MCMC settings
> mcmcOptions <- McmcOptions(iterations=1e4L,
+                             burnin=1e3L,
+                             step=2L)
> ## get samples from the MAP model
> set.seed(634)
> mapSamples <- sampleGlm(models.pima[1L],
+                          mcmc=mcmcOptions,
+                          useOpenMP=FALSE)
```

Taking the linear approximation method
0%-----

Finished MCMC simulation with acceptance ratio 0.868

With the function `McmcOptions`, we have defined an S4 object of MCMC settings, comprising the number of iterations, the length of the burn-in, the thinning step (here save every second iteration), here the acceptance rate was 0.87). Note that you can also get predictive samples for new data points via the `newdata` option of `sampleGlm`. The result `mapSamples` has the following structure:

```
> str(mapSamples)
```

```
List of 5
```

```
$ tbf          : logi FALSE
$ acceptanceRatio: num 0.868
$ logMargLik   :List of 5
..$ estimate           : Named num -266
.. ..- attr(*, "names")= chr "numeratorTerms"
..$ standardError      : num [1, 1] 0.00237
..$ numeratorTerms     : num [1:4500] 0.542 0.657 0.664 0.619 0.559 ...
..$ denominatorTerms  : num [1:4500] 1 1 0.989 1 1 ...
..$ highDensityPointLogUnPosterior: num -266
$ coefficients       : num [1:3, 1:4500] -0.932 4.211 -5.46 -0.827 3.745 ...
..- attr(*, "dimnames")=List of 2
.. ..$ : chr [1:3] "(Intercept)" "glu^1" "skin^-0.5"
.. ..$ : NULL
$ samples           :Formal class 'GlmBayesMfpSamples' [package "glmBfp"] with 8 slots
.. ..@ fitted       : num [1:532, 1:4500] -2.29 2.55 -2.11 1.66 -1.6 ...
.. .. ..- attr(*, "dimnames")=List of 2
.. .. .. ..$ : chr [1:532] "1" "2" "3" "4" ...
.. .. .. ..$ : NULL
.. ..@ predictions  : logi[0 , 0 ]
.. ..@ fixCoefs     :List of 1
.. .. ..$ (Intercept): num [1, 1:4500] -0.932 -0.827 -0.932 -0.936 -1.073 ...
.. .. .. ..- attr(*, "dimnames")=List of 2
.. .. .. .. ..$ : chr "(Intercept)"
.. .. .. .. ..$ : NULL
.. ..@ z            : num [1:4500] 5.32 5.44 7.19 5.08 5.57 ...
.. ..@ bfpCurves   :List of 2
.. .. ..$ glu      : num [1:327, 1:4500] -3.05 -3.02 -3.01 -2.99 -2.96 ...
.. .. .. ..- attr(*, "scaledGrid")= num [1:327, 1] 0.56 0.567 0.57 0.574 0.581 ...
.. .. .. .. ..- attr(*, "dimnames")=List of 2
.. .. .. .. .. ..$ : NULL
```

```

.. .. . . . . .$. : chr "glu"
.. .. . . .- attr(*, "whereObsVals")= int [1:532] 62 318 40 251 113 88 55 313 197 163 ...
.. .. .$. skin: num [1:251, 1:4500] -3.67 -3.47 -3.28 -3.25 -3.11 ...
.. .. . . .- attr(*, "scaledGrid")= num [1:251, 1] 0.7 0.746 0.791 0.8 0.837 ...
.. .. . . .- attr(*, "dimnames")=List of 2
.. .. . . . . $. : NULL
.. .. . . . . $. : chr "skin"
.. .. . . .- attr(*, "whereObsVals")= int [1:532] 67 83 108 115 57 63 76 28 25 95 ...
.. ..@ ucCoefs      : list()
.. ..@ shiftScaleMax: num [1:7, 1:4] 0 0 0 0 1 0 0 10 100 100 ...
.. .. . . .- attr(*, "dimnames")=List of 2
.. .. . . .$. : chr [1:7] "age" "bmi" "bp" "glu" ...
.. .. . . .$. : chr [1:4] "shift" "scale" "maxDegree" "cardPowerset"
.. ..@ nSamples     : int 4500

```

It is a list with the `acceptanceRatio` of the Metropolis-Hastings proposals, an MCMC estimate for the log marginal likelihood including an associated standard error (`logMargLik`), the `coefficients` samples of the model, and an S4 object `samples`. This S4 object includes the `fitted` samples on the linear predictor scale (in our case on the log Odds Ratio scale), possibly `predictions` samples, samples of the intercept (`fixed`), samples of $z = \log(g)$, samples of the fractional polynomial curves (`bfpCurves`), coefficients of uncertain but fixed form covariates (`ucCoefs`), the shifts and scales applied to the original covariates (`shiftScaleMax`) and the number of samples (`nSamples`). You can read more details on the results on the help page by typing `?GlmBayesMfpSamples-class` in R.

If we wanted to get posterior fitted values on the probability scale, we can use the following code:

```

> mapFit <- rowMeans(plogis(mapSamples$samples@fitted))
> head(mapFit)

      1      2      3      4      5      6
0.1017081 0.8967206 0.1096366 0.7864968 0.1794264 0.1412198

```

We can also analyse the MCMC output in greater detail by applying the functions in the `coda` package:

```

> library(coda)
> coefMcmc <- mcmc(data=t(mapSamples$coefficients),
+                 start=mcmcOptions@burnin + 1,
+                 thin=mcmcOptions@step)
> str(coefMcmc)

mcmc [1:4500, 1:3] -0.932 -0.827 -0.932 -0.936 -1.073 ...
- attr(*, "dimnames")=List of 2

```

```

..$ : NULL
..$ : chr [1:3] "(Intercept)" "glu^1" "skin^-0.5"
- attr(*, "mcpars")= num [1:3] 1001 9999 2

> ## standard summary table for the coefficients:
> summary(coefMcmc)

Iterations = 1001:9999
Thinning interval = 2
Number of chains = 1
Sample size per chain = 4500

1. Empirical mean and standard deviation for each variable,
   plus standard error of the mean:

              Mean      SD Naive SE Time-series SE
(Intercept) -0.9298 0.1144 0.001705      0.001812
glu^1        3.8503 0.3963 0.005907      0.006548
skin^-0.5   -4.0626 1.0093 0.015046      0.017079

2. Quantiles for each variable:

              2.5%   25%   50%   75%   97.5%
(Intercept) -1.158 -1.008 -0.929 -0.8514 -0.7113
glu^1        3.085  3.581  3.846  4.0998  4.6522
skin^-0.5   -6.053 -4.736 -4.057 -3.3781 -2.1071

> autocorr(coefMcmc)

, , (Intercept)

              (Intercept)      glu^1      skin^-0.5
Lag 0      1.000000000 -0.198545301  0.22364820
Lag 2      0.060585352 -0.021893505  0.02726208
Lag 10     -0.009556659  0.011005546  0.02048678
Lag 20     -0.008033578  0.001619819 -0.02348717
Lag 100    0.009020242 -0.025143099 -0.01664317

, , glu^1

              (Intercept)      glu^1      skin^-0.5
Lag 0     -0.198545301  1.000000000  0.033350565
Lag 2     -0.015649028  0.102578846 -0.011644669
Lag 10    -0.006802564 -0.007056528 -0.009870305
Lag 20    0.025457688 -0.016328556 -0.004131433

```

```
Lag 100 0.010054085 0.004816710 0.006559927
```

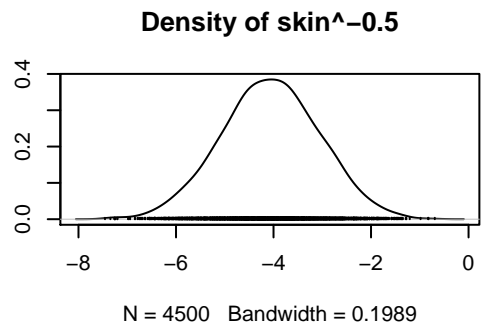
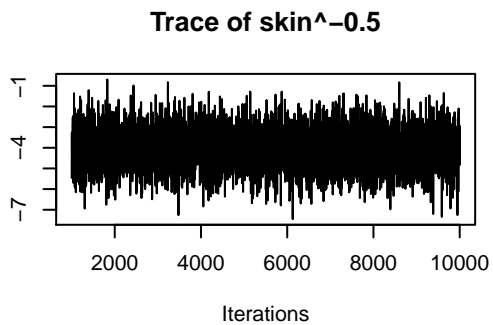
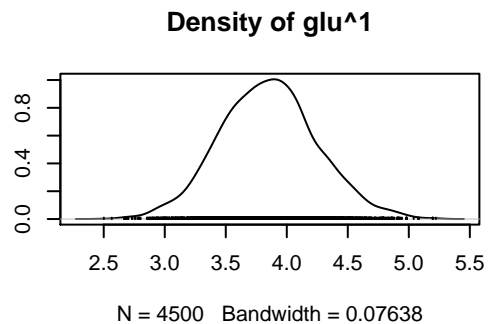
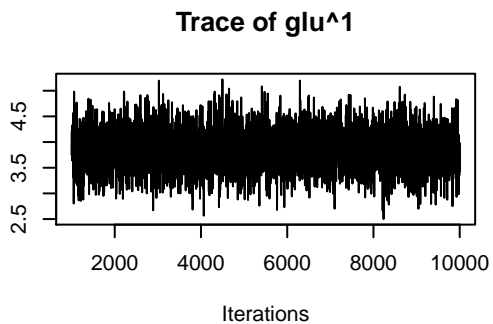
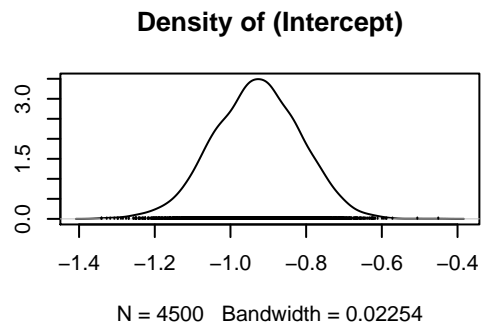
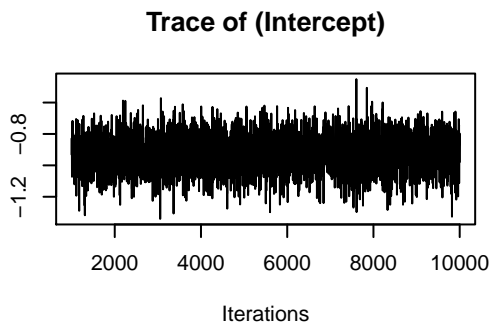
```
, , skin-0.5
```

```
      (Intercept)      glu1      skin-0.5  
Lag 0  0.22364820 0.0333505647 1.0000000000  
Lag 2  0.05308624 0.0007788737 0.0863690518  
Lag 10 0.01053958 -0.0013760574 0.0004120857  
Lag 20 -0.02397061 0.0008594713 -0.0206584865  
Lag 100 -0.02560623 -0.0044481064 -0.0312199771
```

```
> ## etc.
```

```
>
```

```
> plot(coefMcmc)
```

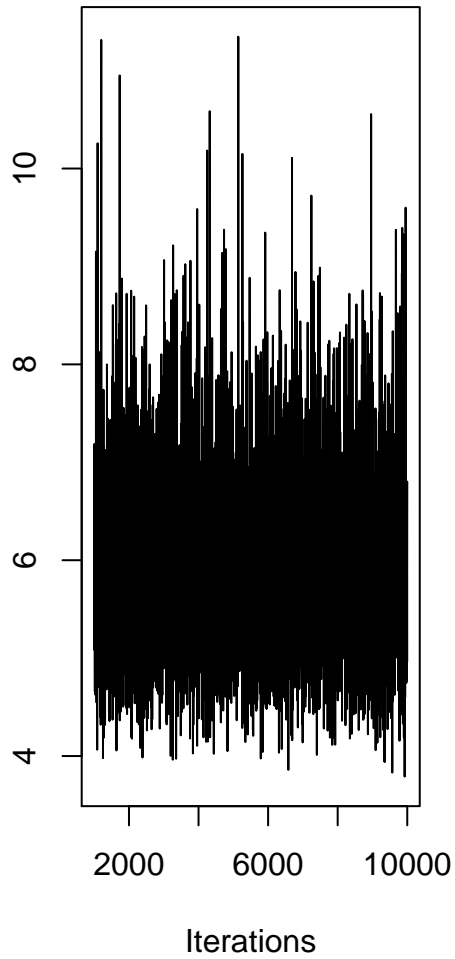



```

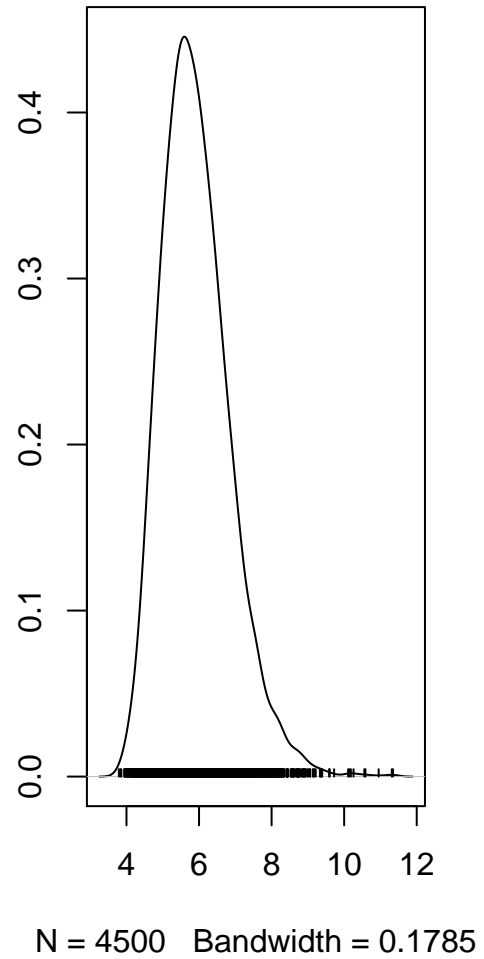
> ## samples of z:
> zMcmc <- mcmc(data=mapSamples$samples@z,
+               start=mcmcOptions@burnin + 1,
+               thin=mcmcOptions@step)
> plot(zMcmc)
> ## etc.

```

Trace of var1

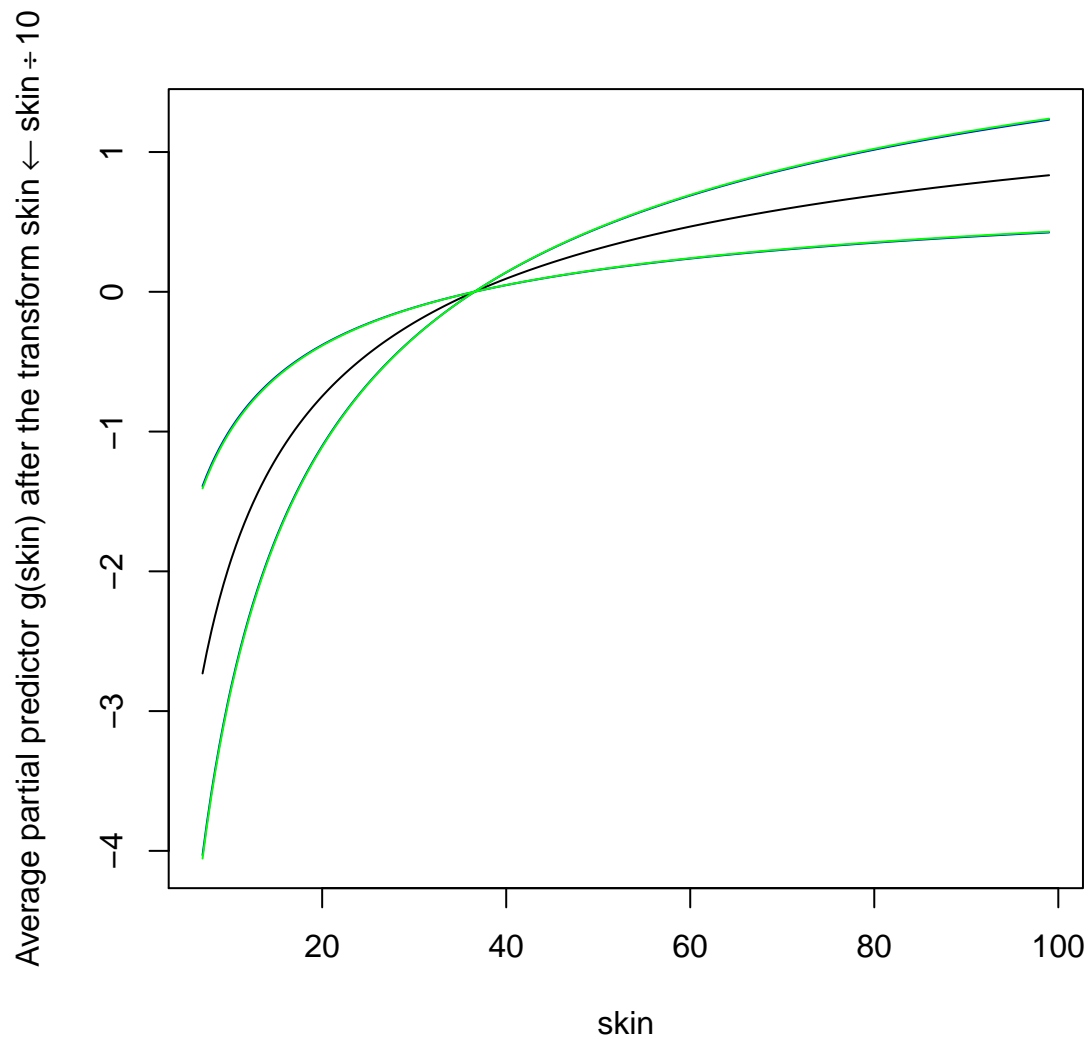


Density of var1



Curve estimates Now we can use the samples to plot the estimated effects of the MAP model covariates, with the `plotCurveEstimate` function. For example:

```
> plotCurveEstimate(termName="skin",  
+                   samples=mapSamples$samples)
```



Model averaging Model averaging works in principle similar to sampling from a single model, but multiple model configurations are supplied and their respective log posterior probabilities. For example, if we wanted to average the top three models found, we would do the following:

```
> set.seed(312)
> bmaSamples <-
+   sampleBma(models.pima[1:3],
+             mcmc=mcmcOptions,
+             useOpenMP=FALSE,
+             nMargLikSamples=1000)
```

```

Starting sampling ...
Now at model 1 ...
Taking the linear approximation method
0%-----
-----
Finished MCMC simulation with acceptance ratio 0.875

Now at model 2 ...
Taking the linear approximation method
0%-----
-----
Finished MCMC simulation with acceptance ratio 0.864

Now at model 3 ...
Taking the linear approximation method
0%-----
-----
Finished MCMC simulation with acceptance ratio 0.886

> ## look at the list element names:
> names(bmaSamples)

[1] "modelData" "samples"

> ## now we can see how close the MCMC estimates ("margLikEstimate")
> ## are to the ILA estimates ("logMargLik") of the log marginal likelihood:
> bmaSamples$modelData[, c("logMargLik", "margLikEstimate")]

  logMargLik margLikEstimate
1 -265.8779      -265.8792
2 -266.2321      -266.2333
3 -270.1569      -270.1514

> ## the "samples" list is again of class "GlmBayesMfpSamples":
> class(bmaSamples$samples)

[1] "GlmBayesMfpSamples"
attr(,"package")
[1] "glmBfp"

```

Then internally, first the models are sampled, and for each sampled model so many samples are drawn as determined by the model frequency in the model average sample. The result is a list with two elements: `modelData` is similar to the `table.pima`, and contains in addition to that the BMA probability and frequency in the sample, the MCMC acceptance ratios (which should be high). On the second element `samples`, which is again of class `GlmBayesMfpSamples`, the above presented functions can again be applied (e.g. `plotCurveEstimate`).

Bibliography

D. Sabanés Bové and L. Held. Hyper- g priors for generalized linear models. 6(3):387–410, 2011. doi: 10.1214/11-BA615. URL <http://projecteuclid.org/euclid.ba/1339616469>.