Heteroscedastic Tobit Regression

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Abstract

The **htobit** package (https://R-Forge.R-project.org/projects/uibk-rprog-2017/) fits tobit regression models with conditional heteroscedasticy using maximum likelihood estimation. A brief overview of the package is provided, along with some illustrations and a replication of results from the **crch** package.

Keywords: heteroscedastic tobit, regression, R.

1. Introduction

The heteroscedastic tobit model assumes an underlying latent Gaussian variable

$$y_i^* \sim \mathcal{N}(\mu_i, \sigma_i^2)$$

which is only observed if positive and zero otherwise: $y_i = \max(0, y_i^*)$. The latent mean μ_i and scale σ_i (latent standard deviation) are linked to two different linear predictors

$$\begin{array}{rcl} \mu_i &=& x_i^\top \beta \\ \log(\sigma_i) &=& z_i^\top \gamma \end{array}$$

where the regressor vectors x_i and z_i can be set up without restrictions, i.e., they can be identical, overlapping or completely different or just including an intercept, etc.

See also Messner, Mayr, and Zeileis (2016) for a more detailed introduction to this model class as well as a better implementation in the package **crch**. The main purpose of **htobit** is to illustrate how to create such a package *from scratch*.

2. Implementation

As usual in many other regression packages for R (R Core Team 2019), the main model fitting function htobit() uses a formula-based interface and returns an (S3) object of class htobit:

```
htobit(formula, data, subset, na.action,
model = TRUE, y = TRUE, x = FALSE,
control = htobit_control(...), ...)
```

Actually, the formula can be a two-part Formula (Zeileis and Croissant 2010), specifying separate sets of regressors x_i and z_i for the location and scale submodels, respectively.

Method	Description							
<pre>print()</pre>	Simple printed display with coefficients							
<pre>summary()</pre>	Standard regression summary; returns summary.htobit object							
	(with print() method)							
coef()	Extract coefficients							
vcov()	Associated covariance matrix							
<pre>predict()</pre>	(Different types of) predictions for new data							
fitted()	Fitted values for observed data							
residuals()	Extract (different types of) residuals							
terms()	Extract terms							
<pre>model.matrix()</pre>	Extract model matrix (or matrices)							
nobs()	Extract number of observations							
logLik()	Extract fitted log-likelihood							
bread()	Extract bread for sandwich covariance							
estfun()	Extract estimating functions (= gradient contributions) for sand-							
	wich covariances							
getSummary()	Extract summary statistics for mtable()							

Table 1: S3 methods provided in **htobit**.

The underlying workhorse function is htobit_fit() which has a matrix interface and returns an unclassed list.

A number of standard S3 methods are provided, see Table 1.

Due to these methods a number of useful utilities work automatically, e.g., AIC(), BIC(), coeftest() (lmtest), lrtest() (lmtest), waldtest() (lmtest), linearHypothesis() (car), mtable() (memisc), Boot() (car), etc.

3. Illustration

To illustrate the package's use in practice, a comparison of several homoscedastic and heteroscedastic tobit regression models is applied to data on budget shares of alcohol and tobacco for 2724 Belgian households (taken from Verbeek 2004). The homoscedastic model from Verbeek (2004) can be replicated by:

```
R> data("AlcoholTobacco", package = "htobit")
R> library("htobit")
R> ma <- htobit(alcohol ~ (age + adults) * log(expenditure) + oldkids + youngkids,
+ data = AlcoholTobacco)
R> summary(ma)
Call:
htobit(formula = alcohol ~ (age + adults) * log(expenditure) +
oldkids + youngkids, data = AlcoholTobacco)
Standardized residuals:
    Min   1Q Median   3Q Max
```

```
-1.0698 -0.4407 -0.1364 0.3934 8.3170
Coefficients (location model):
                         Estimate Std. Error z value Pr(>|z|)
(Intercept)
                       -0.1591533 0.0437782 -3.635 0.000278 ***
                        0.0134938 0.0108824 1.240 0.214989
age
adults
                        0.0291901 0.0169469 1.722 0.084989 .
log(expenditure)
                        0.0126679 0.0032156 3.939 8.17e-05 ***
oldkids
                       -0.0026408 0.0006049 -4.366 1.27e-05 ***
youngkids
                       -0.0038789 0.0023835 -1.627 0.103651
age:log(expenditure)
                       -0.0008093 0.0008006 -1.011 0.312067
adults:log(expenditure) -0.0022484 0.0012232 -1.838 0.066051 .
Coefficients (scale model with log link):
           Estimate Std. Error z value Pr(>|z|)
(Intercept) -3.71236
                       0.01533 -242.1 <2e-16 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Log-likelihood: 4755 on 9 Df
Number of iterations in BFGS optimization: 16
```

This model is now modified in two directions: First, the variables influencing the location parameter are also employed in the scale submodel. Second, because the various coefficients for different numbers of persons in the household do not appear to be very different, a restricted specification for this is used. Using a Wald test for testing linear hypotheses from **car** (Fox and Weisberg 2019) yields

```
R> library("car")
R> linearHypothesis(ma, "oldkids = youngkids")
Linear hypothesis test
Hypothesis:
oldkids - youngkids = 0
Model 1: restricted model
Model 2: alcohol ~ (age + adults) * log(expenditure) + oldkids + youngkids
Df Chisq Pr(>Chisq)
1
2 1 0.2639 0.6075
R> linearHypothesis(ma, "oldkids = adults")
Linear hypothesis test
```

```
Hypothesis:
- adults + oldkids = 0
Model 1: restricted model
Model 2: alcohol ~ (age + adults) * log(expenditure) + oldkids + youngkids
Df Chisq Pr(>Chisq)
1
2 1 3.4994 0.06139 .
---
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Therefore, the following models are considered:

```
R> AlcoholTobacco$persons <- with(AlcoholTobacco, adults + oldkids + youngkids)
R> ma2 <- htobit(alcohol ~ (age + adults) * log(expenditure) + oldkids + youngkids + (age + adults) * log(expenditure) + oldkids + youngkids, data = AlcoholTobacco)
R> ma3 <- htobit(alcohol ~ age + log(expenditure) + persons | age + log(expenditure) + persons, data = AlcoholTobacco)
R> BIC(ma, ma2, ma3)
```

```
df BIC
ma 9 -9439.553
ma2 16 -9735.109
ma3 8 -9777.154
```

The BIC would choose the most parsimonious model but a likelihood ratio test would prefer the unconstrained person coefficients:

```
R> library("lmtest")
R> lrtest(ma, ma2, ma3)
Likelihood ratio test
Model 1: alcohol ~ (age + adults) * log(expenditure) + oldkids + youngkids
Model 2: alcohol ~ (age + adults) * log(expenditure) + oldkids + youngkids |
    (age + adults) * log(expenditure) + oldkids + youngkids
Model 3: alcohol ~ age + log(expenditure) + persons | age + log(expenditure) +
   persons
 #Df LogLik Df
                 Chisq Pr(>Chisq)
1 9 4755.4
2 16 4930.8 7 350.925 < 2.2e-16 ***
   8 4920.2 -8 21.234
3
                        0.006551 **
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

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	htobit		crch			
	location	scale	location	scale	-	
(Intercept)	-0.072^{***}	0.176	-0.072^{***}	0.176	-	
	(0.014)	(0.515)	(0.014)	(0.515)		
age	0.002^{***}	0.064^{**}	* 0.002***	0.064^{**}	*Significance:	$*** \equiv p < 0.001;$
	(0.000)	(0.013)	(0.000)	(0.013)		
$\log(expenditure)$	0.006^{***}	-0.278^{**}	* 0.006***	-0.278^{**}	**	
	(0.001)	(0.038)	(0.001)	(0.038)		
persons	-0.002^{***}	-0.111^{**}	$* -0.002^{***}$	-0.111^{**}	**	
	(0.000)	(0.014)	(0.000)	(0.014)		
		$** \equiv p <$	$< 0.01; * \equiv 1$	p < 0.05		

Table 2:	Replication	of crch	results	using	htobit.
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4. Replication

To assess the reliability of the htobit() implementation, it is benchmarked against the crch() function of (Messner *et al.* 2016), using the same restricted model as above.

Using a model table from **memisc** (Elff 2019) it can be easily seen the results can be replicated using both packages (see Table 2).

R> library("memisc")
R> mtable("htobit" = ma3, "crch" = ca3)

References

- Elff M (2019). memisc: Tools for Management of Survey Data and the Presentation of Analysis Results. R package version 0.99.17.2, URL https://CRAN.R-project.org/package= memisc.
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